

# LARGE DEVIATIONS

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## Historical motivation: Ruin problem

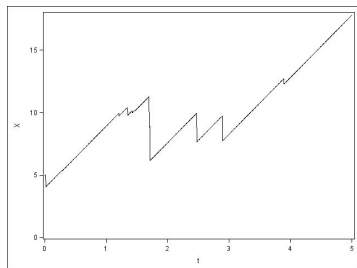
Lundberg (1903). Cramér-Lundberg process:

$$X_t = u + ct - \sum_{i=1}^{N_t} \zeta_i$$

$X_t$  = capital of insurance company at time  $t$ ,

$c$  = premiums income,

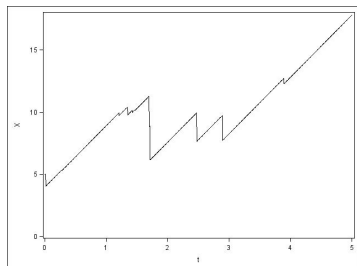
$\{\zeta_i\}$  claims losses.



## Ruin problem (cont.)

Let  $\Psi(u) = \mathbf{P} \{X_t < 0, \text{ for some } t \geq 0 | X_0 = u\}$ .  
Lundberg (1903), Cramér (1930):

$$\Psi(u) \sim Ce^{-Ru} \quad \text{as } u \rightarrow \infty.$$



## Ruin problem (cont.)

Why is this a large deviations problem?

First, exponential decay as  $u \rightarrow \infty$ ; in particular,

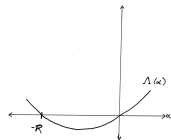
$$\lim_{u \rightarrow \infty} \frac{1}{u} \log \mathbf{P} \{\text{ruin}\} = -R,$$

where  $R$  is the *large deviation rate function*.

Second, proof is based on a change-of-measure argument.

Shifted measure: If  $X_1 \sim \mu$ , then consider

$$\mu_R(dx) := e^{-Rx} \mu(dx).$$



## Statistics: sample means

$X_1, X_2, \dots$  i.i.d.  $\subset \mathbb{R}$ .

Mean  $\mathbf{E}[X_1] = m$ .

Sample mean:

$$\bar{X}_n := \frac{X_1 + \dots + X_n}{n}.$$

Law of large numbers:

$$\mathbf{P}\{|\bar{X}_n - m|\} \rightarrow 0 \quad \text{as } n \rightarrow \infty.$$

How fast?

E.g., what is

$$\lim_{n \rightarrow \infty} \mathbf{P}\{\bar{X}_n > a\} \quad \text{for } a > m?$$

## Statistics: sample means

Central limit approximation:

Let  $S_n = X_1 + \cdots + X_n$ .

Then by the central limit theorem,

$$\bar{X}_n = \frac{S_n}{n} \approx \text{Normal} \left( m, \frac{\sigma^2}{n} \right).$$

How good is this estimate?

## Example

Suppose  $\{X_i\}$  is i.i.d. Bernoulli sequence; thus,

$$X_i = \begin{cases} 1 & \text{w.p. } 1/2, \\ 0 & \text{w.p. } 1/2. \end{cases}$$

Then  $S_n = X_1 + \dots + X_n \sim \text{binomial}$ .

Now according to the central limit theorem, since  $m = \sigma = 1/2$ ,

$$\begin{aligned} \mathbf{P} \left\{ \frac{S_n}{n} > a \right\} &= \mathbf{P} \left\{ \frac{S_n - n \cdot \frac{1}{2}}{\frac{1}{2}\sqrt{n}} > \sqrt{n}(2a - 1) \right\} \\ &\approx 1 - \Phi(\sqrt{n}(2a - 1)). \end{aligned}$$

But if  $a > 1$ , this probability is actually  $= 0$ .

## Cramér's classical estimate

Assume finite moment generating fn.:  $\lambda(\alpha) := \mathbf{E} [e^{\alpha X_i}] < \infty$ .

### Theorem (Cramér, 1938)

For any  $a > m$ ,

$$\mathbf{P} \left\{ \frac{S_n}{n} > a \right\} \sim \frac{c}{\sqrt{n}} e^{-nI(a)} \quad \text{as } n \rightarrow \infty.$$

Shows:

- Exponential decay.
- Decay according to a "rate function"  $I(x)$   
(the convex conjugate of  $\log \lambda$ ).

This rate function will be *different* for every  $\mu$ , where  $X_i \sim \mu$ .

## Empirical measures

$\{X_i\}$  i.i.d.,  $X_i : \Omega \rightarrow \mathbb{S} = \{a_1, \dots, a_d\}$ .

Empirical measure:

$$L_n(a_j) = \frac{1}{n} \sum_{i=1}^n \delta_{X_i}(a_j),$$

that is, the proportion of observed samples which =  $a_j$ .

Note:  $L_n$  is a probability measure on  $\mathbb{S}$ .

Question: How likely will we empirically observe an  $L_n(\cdot)$ , given the true distribution function.

## Empirical measures (cont.)

Sanov (1957): If  $\mathcal{S}$  is a set of empirical measures, then

$$\mathbf{P}\{L_n \in \mathcal{S}\} \approx e^{-nI(\mathcal{S})};$$

more precisely,

$$\lim_{n \rightarrow \infty} \log \mathbf{P}\{L_n \in \mathcal{S}\} = -I(\mathcal{S}).$$

Related problem: Brownian motion in a "tube" (Shilder, 1966).

Consider

$$\mathbf{P}\{B_n(t) \in \Gamma(t) : 0 \leq t \leq 1\}, \quad n \rightarrow \infty,$$

where  $B_n(t) = \frac{1}{\sqrt{n}}B(t)$  and  $B(t)$  is Brownian motion.

## References

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